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Exposure to air pollution and tobacco smoking and their combined effects on depression in six low- and middle-income countries

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Background

Little is known about the joint mental health effects of air pollution and tobacco smoking in low- and middle-income countries.

Aims

To investigate the effects of exposure to ambient fine particulate matter pollution (PM_{2.5}) and smoking and their combined (interactive) effects on depression.

Method

Multilevel logistic regression analysis of baseline data of a prospective cohort study ($n = 41\,785$). The 3-year average concentrations of PM_{2.5} were estimated using US National Aeronautics and Space Administration satellite data, and depression was diagnosed using a standardised questionnaire. Three-level logistic regression models were applied to examine the associations with depression.

Results

The odds ratio (OR) for depression was 1.09 (95% CI 1.01–1.17)

per 10 $\mu\text{g}/\text{m}^3$ increase in ambient PM_{2.5}, and the association remained after adjusting for potential confounding factors (adjusted OR = 1.10, 95% CI 1.02–1.19). Tobacco smoking (smoking status, frequency, duration and amount) was also significantly associated with depression. There appeared to be a synergistic interaction between ambient PM_{2.5} and smoking on depression in the additive model, but the interaction was not statistically significant in the multiplicative model.

Conclusions

Our study suggests that exposure to ambient PM_{2.5} may increase the risk of depression, and smoking may enhance this effect.

Declaration of interest

None.

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Mental illnesses, especially depression, are an increasingly important public health concern, making up three of the ten leading causes of disease burden in low- and middle-income countries.¹ Whereas the cardiopulmonary effects of air pollution have been well established,^{2,3} growing concern has focused on its mental health effects. Several studies have examined the association between exposure to ambient particulate matter with an aerodynamic diameter $\leq 2.5\ \mu\text{m}$ (PM_{2.5}) and depressive symptoms.^{4–6} However, limited studies have been conducted in low- and middle-income countries where higher air pollution levels generally exist. Besides air pollution, tobacco smoking is another important risk factor causing enormous public health burden in low- and middle-income countries. Although the individual effects of both air pollution and smoking have been widely studied, literature on their joint effects has been sparse.^{7,8} One Hong Kong study suggested that smoking could enhance the short-term mortality effects of air pollution;⁷ and Pope *et al* reported an additive interaction between air pollution and smoking on cardiovascular mortality.⁸ Given that both air pollution and smoking have been associated with mental health,⁹ we hypothesise that ambient PM_{2.5} and smoking may have both independent and combined associations with depression. Here we examined the relationship between exposure to ambient PM_{2.5}, smoking and depression among participants in six low- and middle-income countries from the World Health Organization (WHO) Study on global AGEing and adult health (SAGE) Wave 1. We further investigated the potential interaction between PM_{2.5} exposure and tobacco smoking on depression.

Method

Study participants

Details about the SAGE study design and data collection have been reported elsewhere.¹⁰ Briefly, SAGE is a longitudinal cohort study with national samples of adult participants in six low- and middle-income countries: China, Ghana, India, Mexico, Russia and South Africa.¹⁰ This analysis was conducted on cross-sectional data from SAGE Wave 1, which was implemented from 2007 to 2010. The surveys were carried out via a face-to-face household interview using a stratified, multistage, random cluster sampling design. The primary sampling units were stratified by region and location (for example urban/rural) and, within each stratum, enumeration areas were selected. The survey included questions about sociodemographic characteristics, economic situation, behavioural risk factors and health status. This study was approved by the research review boards of the WHO Ethical Review Committee. Informed consent was obtained from each participant.

Depressive symptoms

Depressive symptoms were measured using questions derived from the World Mental Health Survey version of the Composite International Diagnostic Interview. Diagnosis of depression was based on meeting any of the three criteria: (a) self-reported depressive symptoms during the past 12 months; (b) diagnostic criteria for research for depressive episodes; and (c) having received depression-related healthcare services.

PM_{2.5} and smoking

Remote-sensing data were used to estimate the annual concentrations of PM_{2.5} using the method developed by van Donkelaar *et al* to estimate the global distribution of ambient PM_{2.5}.¹¹ The method was based on a combination of observations from the Moderate Resolution Imaging Spectroradiometer and Multiangle Imaging Spectroradiometer instruments from the Terra satellite, and simulations with the GEOS-Chem chemical transport model. The location of the participants' residential community in SAGE was geocoded using Google Earth and used to match the estimated PM_{2.5} concentrations. The word 'community' in the six countries may have different meanings, it is the township or community in China, enumeration area in Ghana and South Africa, village or census enumeration block in India, the Basic Geo-Statistical Area in Mexico, and atenum in Russia. In the main model, the 3-year averaged concentrations of PM_{2.5} before the survey were used as the exposure variable.¹²

In this study, four metrics were used for smoking: smoking status (never or ever), frequency (non-smoking, occasional or daily), duration (never, less or more than 15 years) and amount (never, less or more than eight cigarette equivalents per day).

Covariates

Several covariates were considered, including demographic (age, gender, body mass index (BMI)), socioeconomic (education, annual household income) and lifestyle factors (alcohol consumption, physical activity and cooking-related indoor air pollution). Weight and height were measured to calculate BMI, expressed as weight/height² (kg/m²). Alcohol consumption was categorised into non-drinkers and ever drinkers. To obtain the information about indoor air pollution, participants were asked about their fuel types most frequently used for domestic cooking, and usage of ventilation in indoor cooking spaces. Two fuel types were mainly used: 'clean fuels', including electricity and natural gas, and 'unclean fuels', such as coal, wood, dung and agricultural residues.

The Global Physical Activity Questionnaire was used to measure the intensity, duration, and frequency of physical activity as described elsewhere.¹³ The total time spent in physical activity during the previous week, including the number of days and intensity, was used to generate low, moderate and high categories of physical activity levels.

The country-level indicators were considered, including gross domestic product (GDP) per capita, which was obtained from the Central Intelligence Agency's World Factbook, the percentage of the urban population, per capita healthcare expenditure and the Gini coefficient, a measure of income inequality.

Statistical analysis

A three-level logistic regression model was used with participants being the first-level units, community being the second-level units and country being the third-level units in order to account for the nested structure of the data. In line with previous studies, we calculated the effect estimates for each 10 µg/m³ increase in PM_{2.5}. Multivariate models were conducted to control for some important covariates, including age, gender, BMI, smoking, alcohol drinking, marital status, urbanity, household income, education, domestic fuel type and ventilation. To account for the seasonal pattern of ambient PM_{2.5}, we also included the month of the survey administration in the model as one covariate.

To examine the concentration–response relationship between exposure to ambient PM_{2.5} and depression, we used a quadratic spline smoothing function in the regression model, and visually

examined the shape of the concentration–response relationship. We further examined the possible interaction between PM_{2.5} and smoking (smoking status, frequency, duration and amount) in relation to the risk of depression. We assessed their interactions using both multiplicative and additive models. Multiplicative interaction was examined by adding a product term between PM_{2.5} and smoking into the regression models. Additive-scale interactions were assessed with the synergy index.^{14,15} We classified PM_{2.5} exposure into two categories using the median value as the cut-point. In the additive model, a new variable was constructed to represent the combination of these two variables. Using smoking status (ever, never) as an example, the categories for this variable were as follows: (a) low PM_{2.5} exposure and non-smoker; (b) low PM_{2.5} and smoker; (c) high PM_{2.5} and non-smoker; and (d) high PM_{2.5} and smoker. The synergy index (SI) can be specified as:

$$SI = \frac{OR_{11} - 1}{(OR_{01} - 1) + (OR_{10} - 1)}$$

where OR₁₁ represents the risk in category of high PM_{2.5} and smoker, OR₀₁ is the risk in category of low PM_{2.5} and smoker, and OR₁₀ is the risk in category of high PM_{2.5} and non-smoker. A synergy index greater than one denoted a synergetic interaction, meaning that the joint effects of PM_{2.5} and smoking was greater than the sum of their individual effects. A synergy index smaller than one suggested an antagonistic interaction, indicating that in the presence of two exposures, one factor decreased the effect of the other.¹⁴

Several sensitivity analyses were conducted. First, although country remains as a random effect in the models, additional covariates were further adjusted, including GDP, percentage of residents living in urban areas, per capita healthcare expenditure and Gini coefficient. Second, different PM_{2.5} exposure definitions (1- and 5-year averaged concentration before the survey) were used as the exposure variable. We also conducted sensitivity analyses by dividing the six countries into two groups: one group including only China and India and the other including the other four countries (as China and India had the highest air pollution levels among the six countries). All analyses were conducted using R version 3.2.2. A *P*-value <0.05 was considered statistical significant.

Results

A total of 53 426 participants aged 18 years and older were initially contacted and 47 443 agreed to participate and were recruited in this survey, giving a response rate of 88.8%. Among them, 5581 participants were excluded because of missing information about depression, and the other 77 had missing values for age, gender or other important covariates, the remaining 41 785 participants were included in this analysis (Table 1). The 3-year averaged PM_{2.5} concentration in the six countries was 23.75 µg/m³. South Africa had the lowest level with a mean concentration of 5.97 µg/m³; whereas China and India had the highest concentrations (32.87 and 30.87 µg/m³, respectively). The average age of the participants ranged from 50.0 years in India and 63.1 years in Mexico.

Among the 41 785 participants, 3189 (7.63%) had depressive symptoms. Table 2 shows that participants with depressive symptoms (the depression group) were statistically older than the referent participants (the referent group, 59.8 *v.* 57.8 years), and were exposed to higher PM_{2.5} levels (23.80 *v.* 23.15 µg/m³); the referent group had higher BMI values (24.30 *v.* 23.45 kg/m²). The depression group were more likely to be women, unmarried,

Table 1 Description of population and air pollution characteristics, by country.

Country	Participants	Age, years: mean (s.d.)	PM _{2.5} , µg/m ³			
			Minimum	Mean	Medium	Maximum
China	14 567	60.3 (11.8)	10.66	32.87	31.15	55.53
Ghana	5071	60.1 (14.1)	12.21	17.49	17.44	22.79
India	11 226	50.0 (16.6)	7.86	30.87	27.36	64.08
Mexico	2586	63.1 (14.0)	3.75	10.78	11.14	17.03
Russia	4307	62.3 (13.0)	2.32	6.09	6.17	16.90
South Africa	4028	60.3 (12.3)	1.50	5.97	5.92	20.55
Overall	41 785	57.9 (14.7)	1.50	23.75	19.03	64.08

living in rural areas, without formal education and reporting higher physical activity compared with those without such symptoms. Those in the referent group were more likely to have higher household income and reported consuming alcohol, using clean fuels for domestic cooking and having better ventilation at home compared with the depression group. In particular, we found a significant difference between the two groups in terms of smoking with more individuals in the depression group ever having used tobacco, smoking more frequently, for longer and with a greater tobacco consumption per day.

The concentration–response curve between ambient PM_{2.5} and depression, although not perfect, showed an approximately linear relationship (online Fig. DS1). Table 3 showed significant associations of depression with PM_{2.5} and smoking. The unadjusted odds ratio (OR) was 1.09 (95% CI 1.01–1.17) for each 10 µg/m³ increase in ambient PM_{2.5}. After adjusting for various covariates, the OR was 1.10 (95% CI 1.02–1.19). Smoking was also found to be significantly associated with depression. The odds of depression for ever smokers was 1.11 (95% CI 1.03–1.20) and 1.26 (95% CI 1.15–1.37) in the univariate and multivariate models. There was a significant increasing trend with higher frequency, duration and amount of tobacco smoking; for example, compared with non-smokers, the odds of depression of occasional and daily smokers was 1.33 (95% CI 1.18–1.50) and 1.22 (95% CI 1.10–1.35), respectively; the OR for a smoking duration of fewer than 15 years and more than 15 years was 1.18 (95% CI 1.01–1.38) and 1.21 (95% CI 1.08–1.37), respectively; higher consumption of tobacco was also associated with higher risk of depression (OR = 1.02, 95% CI 0.90–1.16 for those who smoked less than eight cigarette per day, and 1.27, 95% CI 1.13–1.43 for those who smoked more than eight cigarette per day).

Table 4 showed the interaction between PM_{2.5} and smoking on depression using the low PM_{2.5}–non-smoking participants as the reference. The interaction was not significant in multiplicative models. However, the additive models for different smoking metrics showed larger joint effects than the sum of the individual effects. For instance, the individual effect of PM_{2.5} and ever smoked was 1.29 (95% CI 1.09–1.51) and 1.23 (95% CI 1.09–1.39), whereas their joint effect was 1.64 (95% CI 1.38–1.95) with a synergistic index of 1.23. A similar pattern was also observed for smoking frequency, duration and amount with the corresponding synergistic index being 1.32, 1.61 and 1.28, respectively.

The sensitivity analyses suggested that the effect estimates in the main models were robust (online Table DS1). When further adjusting for country-level covariates, we observed a comparable association between PM_{2.5} and depression with an OR of 1.09 (95% CI 1.01–1.18) per 10 µg/m³ increase in PM_{2.5}. Similar results were also observed when using 1- and 5-year averaged PM_{2.5} concentrations as the exposure variable. We observed significant associations in both country groupings, with a smaller effect in China and India (OR = 1.09, 95% CI 1.01–1.18) than in the other four countries (OR = 1.46, 95% CI 1.01–2.12).

Table 2 Comparison of sociodemographic and major risk factors between participants with depressive symptoms and those without such symptoms

Variables	Referent group, <i>n</i> (%)	Depression group, <i>n</i> (%)	<i>P</i> ^a
Age, years: mean (s.d.)	57.8 (14.7)	59.8 (13.9)	0.01
BMI, kg/m ² : mean (s.d.)	24.30 (7.21)	23.45 (7.76)	0.01
PM _{2.5} , µg/m ³ : mean (s.d.)	23.15 (15.56)	23.80 (16.09)	0.03
Gender, <i>n</i> (%)			
Male	16 944 (43.90)	1085 (34.02)	
Female	21 652 (56.10)	2104 (65.98)	0.01
Marital status, <i>n</i> (%)			
Married	27 571 (71.44)	1985 (62.25)	
Unmarried	11 024 (28.56)	1204 (37.75)	0.01
Residence, <i>n</i> (%)			
Urban	18 542 (48.04)	1370 (42.96)	
Rural	20 052 (51.96)	1819 (57.04)	0.01
Education, <i>n</i> (%)			
No formal education	16 946 (43.91)	1975 (61.93)	
Primary school	9498 (24.61)	543 (17.03)	
Middle school	10 014 (25.95)	550 (17.25)	
College or higher	2138 (5.54)	121 (3.79)	0.01
Household income, <i>n</i> (%)			
Low	18 977 (49.17)	1798 (56.38)	
High	19 619 (50.83)	1391 (43.62)	0.01
Drinking, <i>n</i> (%)			
Non-drinker	25 120 (65.08)	2158 (67.67)	
Drinker	13 476 (34.92)	1031 (32.33)	0.01
Physical activity, <i>n</i> (%)			
Low	11 686 (30.28)	1001 (31.39)	
Moderate	8072 (20.91)	593 (18.60)	
High	18 838 (48.81)	1595 (50.02)	0.01
Fuel type, <i>n</i> (%)			
Clean	19 644 (50.90)	1228 (38.51)	
Unclean	18 952 (49.10)	1961 (61.49)	0.01
Ventilation, <i>n</i> (%)			
No	30 780 (79.75)	2737 (85.83)	
Yes	7816 (20.25)	452 (14.17)	0.01
Smoking status, <i>n</i> (%)			
Never	25 316 (65.59)	1923 (60.30)	
Ever	13 280 (34.41)	1266 (39.70)	0.01
Smoking frequency, <i>n</i> (%)			
Never	25 316 (65.59)	1923 (60.32)	
Occasional	4031 (10.46)	413 (12.95)	
Daily	9203 (23.87)	852 (26.73)	0.01
Smoking duration, <i>n</i> (%)			
Never	25 316 (73.98)	1923 (69.32)	
0–15 years	2074 (6.06)	226 (8.15)	
> 15 years	6831 (19.96)	625 (22.53)	0.01
Smoking amount, <i>n</i> (%)			
Never	25 316 (65.59)	1923 (60.30)	
0–8	7717 (19.99)	795 (24.93)	
> 8	5563 (14.41)	471 (14.77)	0.01

BMI, body mass index; PM_{2.5}, fine particulate matter pollution.
a. χ^2 -tests for categorical variables and *t*-tests for continuous variables.

Table 3 Crude and adjusted odds ratio for the prevalence of depressive symptoms associated with long-term exposure to ambient fine particulate matter pollution (PM_{2.5}) and smoking (n = 41 785)

	Odds ratios (95% CI)	
	Crude	Adjusted ^a
Ambient PM _{2.5} , per 10 µg/m ³ increase	1.09 (1.01–1.17)	1.10 (1.02–1.19)
Smoking status		
Never	1.00	1.00
Ever	1.11 (1.03–1.20)	1.26 (1.15–1.37)
Smoking frequency		
Never	1.00	1.00
Occasional	1.19 (1.07–1.32)	1.33 (1.18–1.50)
Daily	1.07 (0.99–1.17)	1.22 (1.10–1.35)
P for trend	0.03	0.01
Smoking duration		
Never	1.00	1.00
0–15 years	1.02 (0.88–1.17)	1.18 (1.01–1.38)
> 15 years	1.13 (1.03–1.24)	1.21 (1.08–1.36)
P for trend	0.01	0.01
Smoking amount		
Never	1.00	1.00
0–8	1.06 (0.94–1.19)	1.02 (0.90–1.16)
> 8	1.05 (0.94–1.16)	1.27 (1.13–1.43)
P for trend	0.29	0.01

a. Adjusted for age, gender, body mass index, smoking (for effect of PM_{2.5}), PM_{2.5} (for smoking), alcohol drinking, marital status, urbanity, household income, education level, domestic fuel type and ventilation.

Discussion

Main findings

This study revealed an association of depression with both PM_{2.5} exposure and tobacco smoking among the adult participants in six low- and middle-income countries. Additionally, using four different smoking metrics of tobacco smoking, the additive models suggested that tobacco smoking may enhance the effect of ambient PM_{2.5} on depression.

Consistent associations were observed in the analyses of the two groups of countries as well as in the combined six countries. However, heterogeneity in the estimated effect magnitude existed with a smaller effect in the group of China and India, although these two countries had the highest concentration levels and the largest ranges of PM_{2.5} levels. The concentration–response relationship between PM_{2.5} and depression might lend some insights into this observation. Although we obtained an approximately linear relationship between ambient PM_{2.5} and depression, we did find a large slope in the countries with low PM_{2.5} concentrations, and a relatively smaller slope in the countries with medium PM_{2.5} concentrations among the six countries. It was also possible that people in both China and India might be more adaptive to the high exposure to PM_{2.5}; another possibility is the differences in the chemical compositions of PM_{2.5}.¹⁶ Further studies exploring these differences are warranted.

Using data from participants in six low- and middle-income countries provided a unique opportunity to tease out the relationship between ambient PM_{2.5} and depression. We recruited a large diverse population, which were representative of the population in the six countries. An additional advantage was the large geographical coverage and diverse spectrum of PM_{2.5} concentrations in the study areas. Furthermore, high-quality survey data were secured through tremendous efforts regarding homogenising the study design, with standardised data collection, and training and certifying data collectors.¹⁰

Table 4 The combined effects between ambient fine particulate matter pollution (PM_{2.5}) exposure and smoking on the risk of depressive symptoms

Category	Odds ratios (95% CI)	
	Crude	Adjusted ^a
PM _{2.5} –smoking status		
Low–non-smoker	1.00	1.00
Low–smoker	1.10 (0.98–1.22)	1.23 (1.09–1.39)
High–non-smoker	1.27 (1.09–1.48)	1.29 (1.09–1.51)
High–smoker	1.42 (1.21–1.67)	1.64 (1.38–1.95)
Synergy index	1.16	1.23
P for multiplicative interaction	0.76	0.67
PM _{2.5} –smoking frequency		
Low–non-smoking	1.00	
Low–occasional	1.25 (1.06–1.47)	1.33 (1.12–1.59)
Low–daily	1.02 (0.90–1.16)	1.18 (1.02–1.35)
High–non-smoking	1.27 (1.09–1.48)	1.29 (1.09–1.51)
High–occasional	1.45 (1.19–1.76)	1.70 (1.38–2.09)
High–daily	1.41 (1.19–1.67)	1.61 (1.34–1.93)
Synergy index ^b	1.42	1.32
P for multiplicative interaction	0.44	0.57
PM _{2.5} –smoking duration		
Low–non-smoking	1.00	1.00
Low–short	1.03 (0.83–1.29)	1.25 (0.98–1.59)
Low–long	1.04 (0.90–1.19)	1.12 (0.96–1.32)
High–non-smoking	1.25 (1.07–1.46)	1.27 (1.08–1.49)
High–short	1.25 (1.00–1.56)	1.45 (1.15–1.84)
High–long	1.50 (1.25–1.79)	1.63 (1.34–1.98)
Synergy index ^b	1.76	1.61
P for multiplicative interaction	0.14	0.22
PM _{2.5} –smoking amount		
Low–non-smoking	1.00	1.00
Low–small	1.13 (0.99–1.28)	1.20 (1.07–1.33)
Low–large	1.04 (0.8–1.22)	1.33 (1.15–1.52)
High–non-smoking	1.27 (1.09–1.48)	1.28 (1.12–1.47)
High–small	1.44 (1.22–1.71)	1.56 (1.34–1.82)
High–large	1.39 (1.15–1.68)	1.78 (1.51–2.11)
Synergy index ^b	1.26	1.28
P for multiplicative interaction	0.72	0.66

a. Adjusted for age, gender, body mass index, alcohol drinking, marital status, urbanity, household income, education level, domestic fuel type and ventilation.
b. The synergy index for PM_{2.5} with smoking frequency, duration and amount was for the combination of high PM_{2.5} exposure with highest levels of smoking variables (smoking daily, more than 15 years and more than eight cigarette equivalents per day).

Comparison with findings from previous studies

Consistent with several previous studies, we found a significant association between smoking and depression.⁹ On the other hand, only a few epidemiological studies have investigated the effects of ambient PM_{2.5} on depression. Both the Nurses' Health Study and a cohort study in Seoul, South Korea showed a positive association between outdoor PM_{2.5} and depressive disorders, which was consistent with our study.^{5,17} One recent study found that blood cadmium levels (one important metal component of PM_{2.5}) were associated with depressive symptoms in elderly people in Seoul.¹⁸ These findings were also consistent with significant depression effects associated with short-term exposure to air pollution.¹⁹ Our findings on positive associations between depression and PM_{2.5} exposure were also supported by literature in the wider context of research on the association between mental health and air pollution. For example, prenatal exposure to air pollution has been associated with increased risk of mental health problems in both human and animal studies.^{20,21} On the other hand, one study from Boston, USA and one large European cohort study did not find significant association between PM_{2.5} and depressive symptoms.^{4,6} The inconsistent results might be the result of differences in the demographic

characteristics of the populations and particle composition in different areas, which has been shown to cause heterogeneity in health effects of particulate pollution.²²

Explanation for our findings

The mechanisms for the association between PM_{2.5} and depression remain largely unknown. One potential mechanism may be DNA methylation, which has been used as a diagnostic biomarker of depression, and has been related to exposure to ambient PM_{2.5}.²³ Other possible pathways included the elicitation of oxidative stress and systemic proinflammatory responses, resulting in systemic inflammation and circulation of inflammatory mediators. The mediators may interact with and activate the cytokine receptors in the brain endothelial cells, generating autoantibodies to cell junction and neural proteins. These autoantibodies have the ability to disrupt the blood–brain barrier, and thus lead to neuroinflammation, brain oxidative stress and neurochemical changes.²⁴ The fine particles have the potential to cross the epithelium, enter the systemic circulation and reach the brain, leading to inflammation and oxidative stress in the nervous system; these inflammatory factors have been closely related to the development of some depressive disorders,²⁵ and observed in human post-mortem brain tissues in individuals chronically exposed to high air pollution levels.²⁶

Our study observed that ambient PM_{2.5} and smoking had a synergistic interaction on depression in the additive model. The underlying mechanism might be the result of the similar pathophysiological pathways of their health effects, such as inflammatory response, interacting with cytokine receptors in brain endothelial cells, causing neuroinflammation and brain oxidative stress.⁸ It is also thought that smoking may decrease the clearance and increase the deposition and retention of fine particles, and thus enhance the health effects.²⁷

Limitations

Our study also had some limitations. Being a cross-sectional study, we cannot identify the temporal relationship between exposure and outcome, and thus, it is impossible for us to establish a causal relationship between PM_{2.5} and depression. We used a 3-year average of satellite-based PM_{2.5} concentration as a surrogate for exposure to ambient PM_{2.5}. These measurements were more likely to reflect the background air pollution levels and may not reflect direct levels of individual exposure. Another limitation was that we had an 11.2% non-participation rate as well as some missing information. However, SAGE aimed to study the overall health status among the populations in the six countries, rather than to specifically determine any impact on health from air pollution. The participants were unaware of the objective of the current study. Thus, differential participation by exposure status was unlikely. It was possible, though, that some individuals with more serious depression may have a higher refusal rate, and thus possibly created a bias in the results. Furthermore, although we considered some important potential confounding factors in data analyses in this study, some residual confounding cannot be totally excluded because of the lack of data on some other important covariates, such as history of depression, noise exposure and neighbourhood environment.

Previous studies have shown that depression was associated with an increased risk of deleterious health behaviours, such as cigarette smoking and alcohol use.⁹ Thus, the history of depression may be a potential confounder that would influence both current depression and smoking behaviour. We will study this research question when the SAGE follow-up data are available. Finally,

we were not able to examine health effects of other ambient air pollutants, such as ozone, on depression, nor control for potential confounding of these other ambient air pollutants because we did not have any exposure data for these other pollutants.

Implications

This study may have some practical implications. The results provide evidence that PM_{2.5} should be considered as one risk factor for depression and more strict air pollution controlling measures should be considered by environmental professionals, especially for China and India where the presence of higher PM_{2.5} exposure existed. However, caution is needed regarding this policy implication, considering that average exposure to air pollution varied greatly among the six study countries, and that socio-economic factors may be strongly associated with exposure to PM_{2.5}. The combined effects of PM_{2.5} and smoking on depression suggested that avoiding high levels of ambient PM_{2.5} may prevent depression among smokers; and more importantly, quitting smoking and reducing the amount smoked are highly recommended to reduce the risk of depression. In summary, the present study provides evidence of a positive association between ambient PM_{2.5} exposure and depression in six low- and middle-income countries, and tobacco smoking may enhance this effect.

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References

- 1 Anwar J, Mpofu E, Matthews LR, Shadoul AF, Brock KE. Reproductive health and access to healthcare facilities: risk factors for depression and anxiety in women with an earthquake experience. *BMC Public Health* 2011; **11**: 523.
- 2 Lin H, Guo Y, Zheng Y, Di Q, Liu T, Xiao J, et al. Long-term effects of ambient pm2.5 on hypertension and blood pressure and attributable risk among older Chinese adults. *Hypertension* 2017; **69**: 806–12.
- 3 Qiu H, Sun S, Tsang H, Wong C-M, Lee RS, Schooling CM, et al. Fine particulate matter exposure and incidence of stroke A cohort study in Hong Kong. *Neurology* 2017; **88**: 1709–17.
- 4 Wang Y, Eliot MN, Koutrakis P, Gryparis A, Schwartz JD, Coull BA, et al. Ambient air pollution and depressive symptoms in older adults: results from the MOBILIZE Boston Study. *Environ Health Perspect* 2014; **122**: 553–8.
- 5 Kim K-N, Lim Y-H, Bae HJ, Kim M, Jung K, Hong Y-C. Long-term fine particulate matter exposure and major depressive disorder in a community-based urban cohort. *Environ Health Perspect* 2016; **124**: 1547–53.
- 6 Zijlema W, Wolf K, Emeny R, Ladwig K, Peters A, Kongsgård H, et al. The association of air pollution and depressed mood in 70,928 individuals from four European cohorts. *Int J Hygiene Environ Health* 2016; **219**: 212–9.
- 7 Wong CM, Ou CQ, Lee NW, Chan KP, Thach TQ, Chau YK, et al. Short-term effects of particulate air pollution on male smokers and never-smokers. *Epidemiol* 2007; **18**: 593–8.
- 8 Pope CA, Burnett RT, Thurston GD, Thun MJ, Calle EE, Krewski D, et al. Cardiovascular mortality and long-term exposure to particulate air pollution epidemiological evidence of general pathophysiological pathways of disease. *Circulation* 2004; **109**: 71–7.
- 9 Boden JM, Fergusson DM, Horwood LJ. Cigarette smoking and depression: tests of causal linkages using a longitudinal birth cohort. *Br J Psychiatry* 2010; **196**: 440–6.
- 10 Kowal P, Chatterji S, Naidoo N, Biritwum R, Fan W, Ridaura RL, et al. Data resource profile: the World Health Organization Study on global AGEing and adult health (SAGE). *Int J Epidemiol* 2012; **41**: 1639–49.
- 11 Van Donkelaar A, Martin RV, Brauer M, Kahn R, Levy R, Verduzco C, et al. Global estimates of ambient fine particulate matter concentrations from satellite-based aerosol optical depth: development and application. *Environ Health Perspect* 2010; **118**: 847.
- 12 Lin H, Guo Y, Zheng Y, Zhao X, Cao Z, Rigdon SE, et al. Exposure to ambient PM_{2.5} associated with overall and domain-specific disability among adults in six low- and middle-income countries. *Environ Int* 2017; **104**: 69–75.
- 13 Lin H, Guo Y, Di Q, Zheng Y, Kowal P, Xiao J, et al. Ambient PM_{2.5} and stroke: effect modifiers and population attributable risk in six low- and middle-income countries. *Stroke* 2017; **48**: 1191–7.
- 14 Andersson T, Alfredsson L, Källberg H, Zdravkovic S, Ahlbom A. Calculating measures of biological interaction. *Eur J Epidemiol* 2005; **20**: 575–9.
- 15 Zhang Z, Xie X, Chen X, Li Y, Lu Y, Mei S, et al. Short-term effects of meteorological factors on hand, foot and mouth disease among children in Shenzhen, China: non-linearity, threshold and interaction. *Sci Total Environ* 2016; **539**: 576–82.
- 16 Lin H, Tao J, Du Y, Liu T, Qian Z, Tian L, et al. Particle size and chemical constituents of ambient particulate pollution associated with cardiovascular mortality in Guangzhou, China. *Environ Pollu* 2016; **208**: 758–66.
- 17 Power MC, Kioumourtzoglou M-A, Hart JE, Okereke OI, Laden F, Weisskopf MG. The relation between past exposure to fine particulate air pollution and prevalent anxiety: observational cohort study. *BMJ* 2015; **350**: h1111.
- 18 Han C, Lim Y-H, Hong Y-C. Does cadmium exposure contribute to depressive symptoms in the elderly population? *Occup Environ Med* 2016; **73**: 355–6.
- 19 Szyzskowicz M, Kousha T, Kingsbury M, Colman I. Air pollution and emergency department visits for depression: a multicity case-crossover study. *Environ Health Insights* 2016; **10**: 155.
- 20 Fonken L, Xu X, Weil ZM, Chen G, Sun Q, Rajagopalan S, et al. Air pollution impairs cognition, provokes depressive-like behaviors and alters hippocampal cytokine expression and morphology. *Mol Psychiatry* 2011; **16**: 987–95.
- 21 Bolton JL, Huff NC, Smith SH, Mason SN, Foster WM, Auten RL, et al. Maternal stress and effects of prenatal air pollution on offspring mental health outcomes in mice. *Environ Health Perspect* 2013; **121**: 1075.
- 22 Dai L, Zanobetti A, Koutrakis P, Schwartz JD. Associations of fine particulate matter species with mortality in the united states: a multicity time-series analysis. *Environ Health Perspect* 2014; **122**: 837–42.
- 23 Madrigano J, Baccarelli A, Mittleman MA, Sparrow D, Spiro A, Vokonas PS, et al. Air pollution and DNA methylation: interaction by psychological factors in the VA Normative Aging Study. *Am J Epidemiol* 2012; **176**: 224–32.
- 24 Hasler G. Pathophysiology of depression: do we have any solid evidence of interest to clinicians? *World Psychiatry* 2010; **9**: 155–61.
- 25 Dantzer R, O'Connor JC, Freund GG, Johnson RW, Kelley KW. From inflammation to sickness and depression: when the immune system subjugates the brain. *Nature Rev Neurosci* 2008; **9**: 46–56.
- 26 Calderon Garciduenas L, Reed W, Maronpot RR, Henriquez-Roldán C, Delgado-Chavez R, Calderon-Garciduenas A, et al. Brain inflammation and Alzheimer's-like pathology in individuals exposed to severe air pollution. *Toxicol Pathol* 2004; **32**: 650–8.
- 27 Möller WB, Kohlhäufel M, Häussinger K, Stahlhofen W, Heyder J. Human alveolar long-term clearance of ferromagnetic iron oxide microparticles in healthy and diseased subjects. *Exp Lung Res* 2001; **27**: 547–68.



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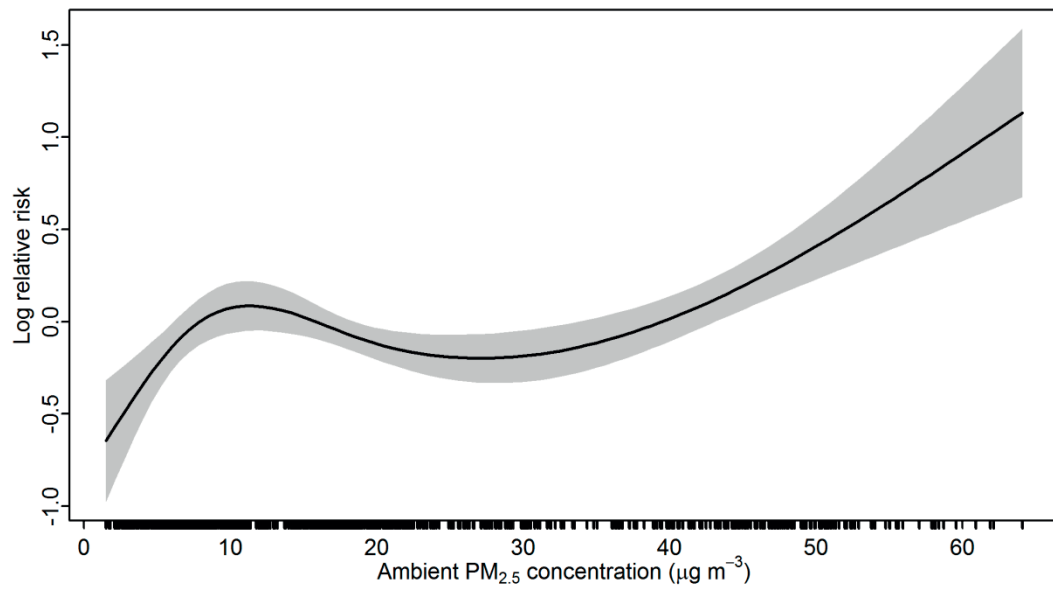


Fig. DS1 The concentration–response curves for ambient PM_{2.5} on depression among adults in six low- and middle-income countries.

Table DS1 Adjusted OR and 95% CI for the prevalence of depression associated with long-term exposure to ambient PM_{2.5} in sensitivity analyses

Models	Odds ratio (OR)	95% CI
Per 10 ug/m ³ increase		
With country-level covariates [#]	1.09	1.01, 1.18
With 1 year PM _{2.5}	1.09	1.01, 1.17
With 5 year PM _{2.5}	1.11	1.03, 1.20
In China and India	1.09	1.01, 1.18
In Ghana, Mexico, Russia and South Africa	1.46	1.01, 2.12

* adjusted for age, sex, BMI, smoking status, alcohol drinking, marital status, urbanity, household income, education level, domestic fuel type and ventilation;

[#] additionally adjusted for country-level covariates, including Gross domestic product per capita, percentage of population living in urban areas, per capita health care expenditure, and Gini coefficient.

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